**Reading Group: An Introduction to PAC-Bayes** Andrew Y. K. Foong, David R. Burt, Javier Antorán 22<sup>nd</sup> April, 2021



Introduction to PAC Generalisation Bounds

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#### Please ask questions!

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In the PAC setting, we view the learning algorithm as choosing a **hypothesis/predictor**  $h \in \mathcal{H}$ .

How shall we choose the hypothesis  $h \in \mathcal{H}$ ?

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• We commonly consider the **0-1 loss**:

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- We want to choose h such that R(h) is minimised.
- However, we don't know D, so we cannot compute R(h).

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- We then compute the **empirical risk** r<sub>S</sub>:

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PAC bounds upper bound true risk  $R(h_S)$  in terms of the empirical risk.

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Hence all randomness comes from sampling S:

- $S, h_S, r_S(h_S), R(h_S)$  are all random variables through S.
- D is not random. R(h) is non-random if h is non-random (in particular, independent of S).

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- PAC bounds constitute a **worst-case analysis**.
- However, very weak assumptions! Typically just i.i.d. assumptions.
- No need to worry about priors or model mismatch!

### The Simplest PAC Bound — Validation

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Want to bound the difference between true and empirical risk.

$$R(h) - r_{5}(h) = \mathbb{E}_{(x,y) \sim D}[\ell((x,y),h)] - \frac{1}{N} \sum_{(x,y) \in S} \ell((x,y),h)$$

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- Concentration inequalities bound deviations of this average.

# Hoeffding's Inequality

#### Theorem 1 (Hoeffding).

Let  $Z_1, \ldots, Z_N$  be i.i.d. random variables bounded in [0,1]. Then for all  $\epsilon > 0$ ,

$$\mathbb{P}\left[\left|\frac{1}{N}\sum_{n=1}^{N}Z_{n}-\mathbb{E}[Z_{n}]\right|>\epsilon\right]\leq 2\exp(-2N\epsilon^{2}).$$

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Probability of a deviation greater than  $\epsilon$  decreases as  $\epsilon$  and N increase.

By writing  $\delta = 2 \exp(-2N\epsilon^2)$ , we get, with probability at least  $1 - \delta$ ,

$$\left|\frac{1}{N}\sum_{n=1}^{N}Z_n-\mathbb{E}[Z_n]\right|\leq \sqrt{\frac{1}{2N}\log\frac{2}{\delta}}$$

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- "Probably"  $\rightarrow$  with probability  $1 \delta$  over dataset *S*, "Approximately"  $\rightarrow$  with a gap term  $\sqrt{\frac{1}{2N} \log \frac{2}{\delta}}$ . But what if we want to choose *h* dependent on *S*?

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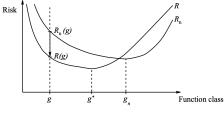


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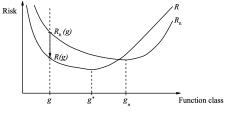


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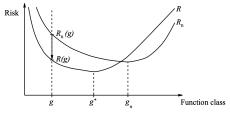


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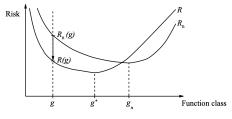


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- Switching order implies that with high probability, the curves  $R_n$  and R are close for all g simultaneously!
- Latter statement more useful: allows us to *choose the hypothesis* depending on S.

# The Union Bound

We seek a statement of the form "with high probability, for all  $h \dots$ "

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$$\mathbb{P}\left[\exists h \in \mathcal{H} : R(h) > r_{\mathcal{S}}(h) + \sqrt{(2N)^{-1}\log(2/\delta)}\right]$$
$$= \mathbb{P}\left[\bigcup_{h \in \mathcal{H}} \left\{S : R(h) > r_{\mathcal{S}}(h) + \sqrt{(2N)^{-1}\log(2/\delta)}\right\}\right]$$
$$\leq \sum_{h \in \mathcal{H}} \mathbb{P}\left[R(h) > r_{\mathcal{S}}(h) + \sqrt{(2N)^{-1}\log(2/\delta)}\right]$$
$$\leq |\mathcal{H}|\delta.$$

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If we set  $\delta' := |\mathcal{H}|\delta$ , we have that with probability at least  $1 - \delta'$ , for all  $h \in \mathcal{H}$  simultaneously,

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- Can we do something more interesting than union bound?
- Yes if we use randomised hypotheses, we can use PAC-Bayes!

PAC-Bayes

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- PAC-Bayes gives bounds of the form  $R_Q \stackrel{1-\delta}{\leq} r_{S,Q} + f(Q, P, N, \delta)$ , for all Q where f depends on how different Q and P are, and usually goes to 0 as  $N \to \infty$ .

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- Warning: The assumptions in PAC-Bayes are different than in Bayes.
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  - No assumption that *P* is related to a generating process/prior beliefs about the data.

## Theorem 2 (Mcallester's Theorem, McAllester, 1999, Maurer Variant).

For any  $\ell \in \{0,1\}$ , D, H and P a probability measure supported on H, for  $N \ge 8$ ,

$$R_{Q} \stackrel{1-\delta}{\leq} r_{S,Q} + \sqrt{\frac{\mathcal{D}_{\mathrm{KL}}\left[Q||P\right] + \log\sqrt{N} + \log\frac{2}{\delta}}{2N}}$$

(1)

for all Q probability measures supported on  $\mathcal{H}$ .

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$$R_{Q} \stackrel{1-\delta}{\leq} r_{S,Q} + \sqrt{\frac{\mathcal{D}_{\mathrm{KL}}\left[Q||P\right] + \log\sqrt{N} + \log\frac{2}{\delta}}{2N}}$$
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- Note if |*H*| < ∞, *P* is uniform and *Q* is a point mass, then
   *D*<sub>KL</sub> [*Q*||*P*] = log |*H*|, in which case this looks a lot like the union bound.

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#### Lemma 3 (Change of Measure).

$$\int \phi(h)q(h)d\lambda - \mathcal{D}_{\mathrm{KL}}\left[Q||P\right] = \log \int e^{\phi(h)}p(h)d\lambda - \mathcal{D}_{\mathrm{KL}}\left[Q||\hat{P}\right] \quad (2)$$
$$\leq \log \int e^{\phi(h)}p(h)d\lambda, \qquad (3)$$

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- This is just the variational 'ELBO'!

#### **Proof of Change of Measure**

We expand out the term  $\mathcal{D}_{\mathrm{KL}}\left[ \mathcal{Q} || \hat{\mathcal{P}} \right]$ :

$$\int \log \frac{q(h)}{\hat{p}(h)} q(h) d\lambda = \int \log \frac{q(h)}{p(h)} \frac{p(h)}{\hat{p}(h)} q(h) d\lambda$$
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The second term on the RHS is,

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#### Theorem 5 (Jensen's inequality).

Let f be a convex function, and suppose  $\mathbb{E}[X], \mathbb{E}[f(X)]$  are finite. Then

 $f(\mathbb{E}[X]) \leq \mathbb{E}[f(X)].$ 

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## Finishing the Proof

It can be shown [Maurer, 2004] that,

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Putting this altogether,

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After rearranging gives,

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- There are also various generalizations e.g. that allow the prior to depend in certain ways on the data S (e.g. Ambroladze et al. [2007]) or that allow for non-i.i.d. data (e.g. Ralaivola et al. [2009]).

## Applications to Neural Networks

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NNs can overfit but in practise don't: why?

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• Finite hypothesis space bound:

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**Bounds are vacuous:** for empirically well-performing models on standard datasets, the generalisation error is bounded by a value greater than 1.

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**Hypothesis:** the complexity of functions found by fitting NN models is much lower than the number of network parameters would suggest.

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r<sub>S,Q</sub> = E[<sup>1</sup>/<sub>N</sub> ∑<sub>(x,y)∈S</sub> 1{h(x) ≠ y}] is not differentiable! Use convex surrogate upper bound:

$$r_{S,Q} \leq r_{\overline{S},Q} = \mathbb{E}_{h \sim Q}\left[\frac{1}{N\sqrt{2}}\sum_{(x,y) \in S} \log(1 + \exp(-h(x)y))\right]$$

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- Choose P = N(w; 0, λl). λ is chosen from a predefined set using a union bound.

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**4** Estimate bound using original  $r_{S,Q}$  loss and samples from Q. Intuition:

- Local optima in flat regions have a smaller description length
- This approach is very similar to Bayes by Backprop: we are approximately optimising a lower bound on the marginal likelihood.

Experiment	T-600	T-1200	$T-300^{2}$	$T-600^{2}$	$T-1200^{2}$	$T-600^{3}$	R-600
Train error	0.001	0.002	0.000	0.000	0.000	0.000	0.007
Test error	0.018	0.018	0.015	0.016	0.015	0.013	0.508
PAC-Bayes bound	0.161	0.179	0.170	0.186	0.223	0.201	1.352
KL divergence	5144	5977	5791	6534	8558	7861	201131
# parameters	471k	943k	326k	832k	2384k	1193k	472k
VC dimension	26m	56m	26m	66m	187m	121m	26m

Table 1: Results for experiments on binary class variant of MNIST. SGD is either trained on (T) true labels or (R) random labels. The network architecture is expressed as  $N^L$ , indicating L hidden layers with N nodes each. Errors are classification error. The reported VC dimension is the best known upper bound (in millions) for ReLU networks. The SNN error rates are tight upper bounds (see text for details). The PAC-Bayes bounds upper bound the test error with probability 0.965.

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#### Takeaways:

- Bounds are less than 1 when models perform well
- Bounds can warn us when our model will not generalise

## Scaling to Imagenet using NN compression

**Observation:** 

- The KL divergence  $\mathcal{D}_{\mathrm{KL}}[Q||P]$  can be seen as the expected number of bits needed to encode a message sampled from Q using a coding scheme optimal for P.

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Zhou et al. [2019] leverage this interpretation to derive bounds for large networks after **pruning and quantization**.

On Imagenet, they obtain a bound of 96.5% while the validation error is 35%. (Non-vacuous!)

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Using Jensen's we can see:

$$\underbrace{\mathbb{E}_{(x,y)\sim D}[-\log \mathbb{E}_{h\sim Q}[p(y|x,h)]}_{CE_Q} \leq \underbrace{\mathbb{E}_{h\sim Q}[\mathbb{E}_{(x,y)\sim D}[-\log p(y|x,h)]]}_{R_Q}$$

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Minimising the above bound seems like it could be a good idea. Does the optima of  $R_Q$  also minimise  $CE_Q$ ?

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Otherwise we are learning under model misspecification.

Recall:

$$\underbrace{\mathbb{E}_{(x,y)\sim D}[-\log \mathbb{E}_{h\sim Q}[p(y|x,h)]}_{CE_Q} \leq \underbrace{\mathbb{E}_{h\sim Q}[\mathbb{E}_{(x,y)\sim D}[-\log p(y|x,h)]]}_{R_Q}$$

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This will only be a minimiser of  $CE_Q$  if:

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for all Q.

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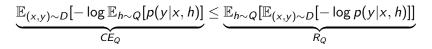
$$\mathcal{D}_{\mathrm{KL}}\left[D_{y|x}||p(y|x,h^{\mathsf{ML}})
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Recall:

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$$CE_Q \leq R_Q - V_Q \leq R_Q$$

where  $V_Q$  is a variance encouraging term.

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 $V_Q$  takes positive values for posteriors different than a delta. It reduces to 0 otherwise (perfect model specification).

### Second order PAC-Bayes: Illustration

We can add this new term to our PAC-Bayes bound:

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#### Second order PAC-Bayes: Illustration

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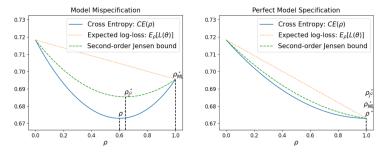


Figure 2: First-Order vs Second-Order Jensen Bounds. See Appendix B for full details.

### Second order PAC-Bayes: Misspecified noise model

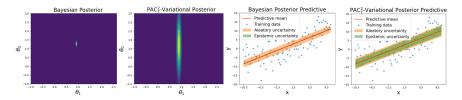


Figure 1: The exact Bayesian posterior and our new proposed (PAC<sup>2</sup><sub>T</sub>-Variational) posterior, and their respective posterior predictive distributions, for a linear regression model with a misspecified constant noise term (the data noise is higher than the linear model's noise). The Bayesian posterior concentrates around the best single linear model, while our method estimates a posterior which introduces high variance in the intercept parameter  $\theta_0$  to induce a posterior predictive distribution with higher noise that better fits the data distribution (see Appendix C.2] for details).

The new variance term is able to increase disagreement among hypothesis, increasing predictive variance.

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